

Importance Sampling (IS)

Problem: estimate $\mathbb{E}_{x \sim p} [f(x)]$, where $p \in \Delta(\mathcal{X})$ and $f: \mathcal{X} \rightarrow \mathbb{R}$.

MC: draw $x_1, \dots, x_n \stackrel{iid}{\sim} p$, estim: $\frac{1}{n} \sum_{i=1}^n f(x_i)$. if f is bounded. Hoeffding's ineq applies.
 (in short, will write: $x \sim p$. $f(x)$ as the estimator).

→ Example: MC policy eval. $J(\pi) := \mathbb{E}_{\pi} \left[\sum_{t=1}^H \gamma^{t-1} r_t \right]$.
 $x \leftrightarrow \tau = (s_1, a_1, r_1, s_2, a_2, r_2, \dots, s_H, a_H, r_H)$
 $\phi \leftrightarrow \tau \sim \pi$. (from obs).
 $f \leftrightarrow \tau \mapsto \sum_{t=1}^H \gamma^{t-1} r_t$.

IS: no access to $x \sim p$. but can sample $x \sim q$, where $q \in \Delta(\mathcal{X})$.

if: $\forall x$ where $p(x) > 0$, we have $q(x) > 0$. then: \leftarrow

$x \sim q$, $\frac{p(x)}{q(x)} f(x)$ \leftarrow IS estimator. IPS, IW.
 imp. weights/ratio, density ratios, etc.

IS: design q .
 IPS, IW: $\frac{p(x)}{q(x)}$.

Claim: IS is unbiased

assume \mathcal{X} is finite.

Proof: $\mathbb{E}_{x \sim q} \left[\frac{p(x)}{q(x)} f(x) \right] = \sum_{x \in \mathcal{X}} q(x) \cdot \left(\frac{p(x)}{q(x)} f(x) \right) = \mathbb{E}_{x \sim p} [f(x)]$.

ideally: $\frac{p(x)}{q(x)}$ should be small. | if $\max_x \frac{p(x)}{q(x)} = C$. $f(v) \in [-B, B]$.

$\downarrow [-CB, CB]$.

Fact: $\mathbb{E}_{x \sim q} \left[\frac{p(x)}{q(x)} \right] = 1$

$\| \frac{p}{q} \|_{\infty} \leftarrow$

$\mathbb{E}_{x \sim q} \left[\frac{p(x)^2}{q(x)^2} \right] \leftarrow$

Example: OPE in contextual bandit

S : contexts, A : actions, $R: S \times A \rightarrow \Delta([0, 1])$.

let $d_0 \in \Delta(S)$ be the ctx dist.

behavior/logging policy.

Have data: $\{(s, a, r)\}$: $s \sim d_0$, $a \sim \pi_b(\cdot | s)$, $r \sim R(\cdot | s, a)$.

Goal: estimate $J(\pi) := \mathbb{E}[r | \pi]$. ↙ target/eval policy.

IPS: $\frac{\pi(a|s)}{\pi_b(a|s)} \cdot r := \underline{p} \cdot r$.

Proof of unbiasedness: Let $(s, a, r) \sim q \Leftrightarrow s \sim d_0, \boxed{a \sim \pi_b}, r \sim R(\text{data})$.
 Let $(s, a, r) \sim p \Leftrightarrow s \sim d_0, \boxed{a \sim \pi}, r \sim R$.

$$\begin{aligned} J(\pi) &= \mathbb{E}_{(s, a, r) \sim p} [r] = \mathbb{E}_{(s, a, r) \sim q}^{\text{"data"}} \left[\frac{p(s, a, r)}{q(s, a, r)} \cdot r \right] \\ &= \mathbb{E}_{(s, a, r) \sim q} \left[\frac{p(s) \cdot p(a|s) \cdot p(r|s, a)}{q(s) \cdot q(a|s) \cdot q(r|s, a)} \cdot r \right] \\ &= \mathbb{E}_{(s, a, r) \sim q} \left[\frac{d_0(s) \cdot \pi(a|s) \cdot R(r|s, a)}{d_0(s) \cdot \pi_b(a|s) \cdot R(r|s, a)} \cdot r \right] = \mathbb{E}_{(s, a, r) \sim q} \left[\frac{\pi(a|s)}{\pi_b(a|s)} r \right] \end{aligned}$$

Variance of IS: Consider special case where

- π is deterministic. $a = \pi(s)$ let $K = |A|$.
- π_b is unif over A . (or $\pi_b \sim U$). ←

IS: $a \sim U, p, r$, where $p = \frac{\pi(a|s)}{\pi_b(a|s)} = \frac{\mathbb{I}(a = \pi(s))}{1/K}$.

Let's further assume that \underline{r} is const (ind. of s, a , has no randomness). γ_0

$$\begin{aligned} \text{Var}[p\underline{r}] &= \gamma_0^2 \text{Var}[p] \\ &= \gamma_0^2 \cdot (E[p^2] - (E[p])^2) = \gamma_0^2 (E[p^2] - 1) \\ &= \gamma_0^2 \left(\mathbb{E} \left[\frac{\mathbb{I}(a = \pi(s))}{1/K^2} \right] - 1 \right) \\ &= \gamma_0^2 \left(K \mathbb{E} \left[\frac{\mathbb{I}(a = \pi(s))}{1/K} \right] - 1 \right) = \boxed{\gamma_0^2 (K - 1)} \end{aligned}$$

IS: $\{(s_i, a_i, r_i)\}_{i=1}^n$

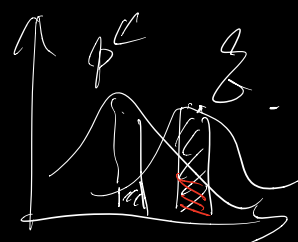
$$\frac{1}{n} \sum_{i=1}^n \frac{\mathbb{I}(a_i = \pi(s_i))}{1/K} r_i = \frac{1}{n/K} \sum_{i: a_i = \pi(s_i)} r_i$$

Can we address this?

- Improvement 1: WIS. (self-normalized IPS).

In the above spec. case:

$$\frac{\sum_{i: a_i = \pi(s_i)} r_i}{|\{i: a_i = \pi(s_i)\}|}$$



General case: $\frac{\sum_{i=1}^n p_i \cdot r_i}{\sum_{i=1}^n p_i} \xrightarrow{\text{expe}} n$

(biased but consistent).

Improvement 2: DR (doubly robust)

(unbiased).

In the spec case: $\frac{p \cdot r}{p} \Rightarrow \hat{r}_0 + p(r - \hat{r}_0)$

General case: $\hat{R} : S \times A \rightarrow \mathbb{R}$
arbitrary function.

$$\begin{aligned} E[p \cdot r_0] &= r_0 \cdot E[p] \\ &= r_0 \end{aligned}$$

$$DR = E_{a \sim \pi} [\hat{R}(s, a)] + p \cdot (r - \hat{R}(s, a))$$

"central variate"

[Dudik, Li, Langford]

Why DR if we have good \hat{R} ?

→ DR has low var. if $\hat{R} \approx R$.

→ DR is always unbiased even if \hat{R} is poor.

→ IS is a special case of DR: $\hat{R} \equiv 0$.

regress

$(s, a) \rightarrow r$
on separate data
to fit \hat{R} .

Applications to RL