# Concentration Inequalities and Union Bound 

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This note introduces the basics of concentration inequalities and examples of its applications (often with union bound), which will be useful for the rest of this course.

## 1 Hoeffding's Inequality

Theorem 1. Let $X_{1}, \ldots, X_{n}$ be independent random variables on $\mathbb{R}$ such that $X_{i}$ is bounded in the interval $\left[a_{i}, b_{i}\right]$. Let $S_{n}=\sum_{i=1}^{n} X_{i}$. Then for all $t>0$,

$$
\begin{array}{r}
\operatorname{Pr}\left[S_{n}-\mathbb{E}\left[S_{n}\right] \geq t\right] \leq e^{-2 t^{2} / \sum_{i=1}^{n}\left(b_{i}-a_{i}\right)^{2}}, \\
\operatorname{Pr}\left[S_{n}-\mathbb{E}\left[S_{n}\right] \leq-t\right] \leq e^{-2 t^{2} / \sum_{i=1}^{n}\left(b_{i}-a_{i}\right)^{2}} \tag{2}
\end{array}
$$

## Remarks:

- By union bound, we have $\operatorname{Pr}\left[\left|S_{n}-\mathbb{E}\left[S_{n}\right]\right| \geq t\right] \leq 2 e^{-2 t^{2} / \sum_{i=1}^{n}\left(b_{i}-a_{i}\right)^{2}}$.
- We often care about the convergence of the empirical mean to the true average, so we can devide $S_{n}$ by $n: \operatorname{Pr}\left[\left|\frac{S_{n}}{n}-\frac{\mathbb{E}\left[S_{n}\right]}{n}\right| \geq t\right] \leq 2 e^{-2 n^{2} t^{2} / \sum_{i=1}^{n}\left(b_{i}-a_{i}\right)^{2}}$.
- A useful rephrase of the result when all variables share the same support $[a, b]$ : with probability at least $1-\delta,\left|\frac{S_{n}}{n}-\frac{\mathbb{E}\left[S_{n}\right]}{n}\right| \leq(b-a) \sqrt{\frac{1}{2 n} \ln \frac{2}{\delta}}$.
- $X_{1}, \ldots, X_{n}$ are not necessarily identically distributed; they just have to be independent.
- The number of variables, $n$, is a constant in the theorem statement. When $n$ is a random variable itself, for Hoeffding's inequality to apply, $n$ cannot depend on the realization of $X_{1}, \ldots, X_{n}$. Example: Consider the following Markov chain:


Say we start at $s_{1}$ and sample a path of length $T$ ( $T$ is a constant). Let $n$ be the number of times we visit $s_{1}$, and we can use the transitions from $s_{1}$ to estimate $p$.

1. Can we directly apply Hoeffding's inequality here with $n$ as the number of coin tosses? If you want to derive a concentration bound for this problem, look up Azuma's inequality.
2. What if we sample a path until we visit $s_{1} N$ times for some constant $N$ ? Can we apply Hoeffding's inequality with $N$ as the number of random variables?

## 2 Multi-Armed Bandits (MAB)

### 2.1 Formulation

A MAB problem is specified by $K$ distributions over $\mathbb{R},\left\{R_{i}\right\}_{i=1}^{K}$. Each $R_{i}$ has bounded supported $[0,1]$ and mean $\mu_{i}$. Let $\mu^{\star}=\max _{i \in[K]} \mu_{i}$. For round $t=1,2, \ldots, T$, the learner

1. Chooses arm $i_{t} \in[K]$.
2. Receives reward $r_{t} \sim R_{i_{t}}$.

A popular objective for MAB is the pseudo-regret, which poses the exploration-exploitation challenge:

$$
\operatorname{Regret}_{T}=\sum_{t=1}^{T}\left(\mu^{\star}-\mu_{i_{t}}\right)
$$

Another important objective is the simple regret:

$$
\mu^{\star}-\mu_{\hat{i}},
$$

where $\hat{i}$ is the arm that the learner picks after $T$ rounds of interactions. This poses the "pure exploration" challenge, since all it matters is to make a good final guess and the regret incurred within the $T$ rounds does not matter. A related objective is called Best-Arm Identification, which asks whether $\hat{i} \in \arg \max _{i \in[K]} \mu_{i} ;$ Best-Arm Identification results often require additional gap conditions.

### 2.2 Uniform sampling

We consider the simplest algorithm that chooses each arm the same number of times, and after $T$ rounds selects the arm with the highest empirical mean. For simplicity let's assume that $T / K$ is an integer. We will prove a high-probability bound on the simple regret. The analysis gives an example of the application of Hoeffiding's inequlaity to a learning problem; the algorithm itself is likely to be suboptimal.

For simplicity let's assume that $T / K$ is an integer. After $T$ rounds, each arm is chosen $T / K$ times, and let $\hat{\mu}_{i}$ be the empirical average reward associated with arm $i$. By Hoeffding's inequality, we have:

$$
\operatorname{Pr}\left[\left|\hat{\mu}_{i}-\mu_{i}\right| \geq \epsilon\right] \leq 2 e^{-2 T \epsilon^{2} / K}
$$

Now we want accurate estimation for all arms simultaneously. That is, we want to bound the probability of the event that any $\hat{\mu}_{i}$ deviating from $\mu_{i}$ too much. This is where union bound is useful:

$$
\begin{aligned}
& \operatorname{Pr}\left[\bigcup_{i=1}^{K}\left\{\left|\hat{\mu}_{i}-\mu_{i}\right| \geq \epsilon\right\}\right] \quad \text { (the event that estimation is } \epsilon \text {-inaccurate for at least } 1 \text { arm) } \\
\leq & \sum_{i=1}^{K} \operatorname{Pr}\left[\left|\hat{\mu}_{i}-\mu_{i}\right| \geq \epsilon\right] \leq 2 K e^{-2 T \epsilon^{2} / K .} \quad \text { (union bound, then Hoeffding's inequality) }
\end{aligned}
$$

To rephrase this result: with probability at least $1-\delta,\left|\hat{\mu}_{i}-\mu_{i}\right| \leq \sqrt{\frac{K}{2 T} \ln \frac{2 K}{\delta}}$ holds for all $i$ simultaneously.

Finally, we use the estimation error to bound the decision loss: recall that $\hat{i}=\arg \max _{i \in[K]} \hat{\mu}_{i}$, and let $i^{\star}=\arg \max _{i \in[K]} \mu_{i}$.

$$
\begin{aligned}
\mu^{\star}-\mu_{\hat{i}} & =\mu_{i^{\star}}-\hat{\mu}_{i^{\star}}+\hat{\mu}_{i^{\star}}-\mu_{\hat{i}} \\
& \leq \mu_{i^{\star}}-\hat{\mu}_{i^{\star}}+\hat{\mu}_{\hat{i}}-\mu_{\hat{i}} \leq 2 \sqrt{\frac{K}{2 T} \ln \frac{2 K}{\delta}}
\end{aligned}
$$

We can rephrase this result as a sample complexity statement: in order to guarantee that $\mu^{\star}-\mu_{\hat{i}} \leq \epsilon$ with probablity at least $1-\delta$, we need $T=O\left(\frac{K}{\epsilon^{2}} \ln \frac{K}{\delta}\right)$.

### 2.3 Lower bound

The linear dependence of the sample complexity on $K$ makes a lot of sense, as to choose a arm with high reward we have to try each arm at least once. Below we will see how to mathematically formalize this idea and prove a lower bound on the sample complexity of MAB.

Theorem 2. For any $K \geq 2, \epsilon \leq \sqrt{1 / 8}$, and any $M A B$ algorithm, there exists an $M A B$ instance where $\mu^{\star}$ is $\epsilon$ better than other arms, yet the algorithm identifies the best arm with no more than $2 / 3$ probability unless $T \geq \frac{K}{72 \epsilon^{2}}$.

The theorem itself is stated as a best-arm identification lower bound, but it is also a lower bound for simple regret minimization. This is because all arms except the best one is $\epsilon$ worse than $\mu^{\star}$, so missing the optimal arm means a simple regret of at least $\epsilon$.

See the proof in [1] (Theorem 2); the technique is due to [2] and can be also used to prove the lower bound on the regret of MAB.

## 3 Generalization Bounds for Supervised Learning

Consider a simple supervised learning setting: let $\mathcal{X}$ be the feature space and $\mathcal{Y}$ be the label space; in this example we consider classification so $\mathcal{Y}=\{0,1\}$. Let $P_{X, Y}$ be a distribution over $\mathcal{X} \times \mathcal{Y}$, and we are given a dataset $\left\{\left(X_{i}, Y_{i}\right)\right\}_{i=1}^{n}$ with each $\left(X_{i}, Y_{i}\right)$ drawn i.i.d. from $P_{X, Y}$. Let $\mathcal{F}: \mathcal{X} \rightarrow \mathcal{Y}$ be a finite hypothesis class. The classifier in $\mathcal{F}$ that minimizes the classification error is:

$$
f^{\star}:=\underset{f \in \mathcal{F}}{\arg \min } \mathbb{E}[\mathbb{I}[f(X) \neq Y]],
$$

where $\mathbb{E}[\cdot]$ is w.r.t. $P_{X, Y}$. Given only a finite sample, one natural thing to do is empirical risk minimization, i.e., find the classifer that has the lowest training error rate on data:

$$
\hat{f}=\underset{f \in \mathcal{F}}{\arg \min } \widehat{\mathbb{E}}[\mathbb{I}[f(X) \neq Y]]:=\frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\left[f\left(X_{i}\right) \neq Y_{i}\right] .
$$

The question is, can we give any guarantee to how good the learned classifier $\hat{f}$ is compared to the optimal one $f^{\star}$, as a function of $n$ ? In other words, we want to bound

$$
\left.\mathbb{E}[\mathbb{I}[\hat{f}(X) \neq Y)]]-\mathbb{E}\left[\mathbb{I}\left[f^{\star}(X) \neq Y\right)\right]\right]
$$

We provide the analysis below, which mainly uses Hoeffding's and union bound. First of all,

$$
\begin{align*}
& \left.\mathbb{E}[\mathbb{I}[\hat{f}(X) \neq Y)]]-\mathbb{E}\left[\mathbb{I}\left[f^{\star}(X) \neq Y\right)\right]\right] \\
\leq & \left.\left.\mathbb{E}[\mathbb{I}[\hat{f}(X) \neq Y)]]-\widehat{\mathbb{E}}[\mathbb{I}[\hat{f}(X) \neq Y)]]+\widehat{\mathbb{E}}\left[\mathbb{I}\left[f^{\star}(X) \neq Y\right)\right]\right]-\mathbb{E}\left[\mathbb{I}\left[f^{\star}(X) \neq Y\right)\right]\right] \quad(\hat{f} \text { is optimal w.r.t. } \widehat{\mathbb{E}}) \\
\leq & \left.\left.2 \cdot \max _{f \in \mathcal{F}} \mid \mathbb{E}[\mathbb{I}[f(X) \neq Y)]\right]-\widehat{\mathbb{E}}[\mathbb{I}[f(X) \neq Y)]\right] \mid \tag{3}
\end{align*}
$$

It then suffices to bound $\left.\left.\max _{f \in \mathcal{F}} \mid \mathbb{E}[\mathbb{I}[f(X) \neq Y)]\right]-\widehat{\mathbb{E}}[\mathbb{I}[f(X) \neq Y)]\right] \mid$, which is often called a uniform deviation bound. The key is to realize that, for any fixed $f \in \mathcal{F}, \widehat{\mathbb{E}}[\mathbb{I}[f(X) \neq Y]]$ is the average of i.i.d. random variables $\mathbb{I}\left[f\left(X_{i}\right) \neq Y_{i}\right]$ bounded in $[0,1]$, whose true expectation is precisely $\mathbb{E}[\mathbb{I}[f(X) \neq$ $Y]]$. Applying Hoeffding's, for a fixed $f \in \mathcal{F}$, with probability at least $1-\delta$, we have

$$
\left\lvert\, \widehat{\mathbb{E}}\left[\mathbb{I}[f(X) \neq Y]-\mathbb{E}\left[\mathbb{I}[f(X) \neq Y] \left\lvert\, \leq \sqrt{\frac{1}{2 n} \ln \frac{2}{\delta}}\right.\right.\right.\right.
$$

Union bounding over $\mathcal{F}$ and plugging into Eq.(4),

$$
\begin{equation*}
\left.\mathbb{E}[\mathbb{I}[\hat{f}(X) \neq Y)]]-\mathbb{E}\left[\mathbb{I}\left[f^{\star}(X) \neq Y\right)\right]\right] \leq \sqrt{\frac{2}{n} \ln \frac{2|\mathcal{F}|}{\delta}} \tag{4}
\end{equation*}
$$

## References

[1] Akshay Krishnamurthy, Alekh Agarwal, and John Langford. PAC reinforcement learning with rich observations. In Advances in Neural Information Processing Systems, pages 1840-1848, 2016.
[2] Peter Auer, Nicolò Cesa-Bianchi, and Paul Fischer. Finite-time analysis of the multiarmed bandit problem. Machine learning, 47(2-3):235-256, 2002.

